

## Sampling-reconstruction procedure of discrete Markov processes with continuous time

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### Summary

At the first time the statistical description of the Sampling-Reconstruction Procedure (SRP) of Discrete Markov Processes (Markov chains) with continuous time and with an arbitrary number of states is given. The mathematical models of Markov chains with continuous time are intensively used in the description of some real stochastic processes with jumps (in control systems and radio engineering), for instance, impulse noise [1, 2]. This is the reason that it is necessary to know: how to sample, how to reconstruct and how to calculate the reconstruction errors of such processes. (Jumps can be occurred in continuous time moments.) So, the usual method of the SRP investigation of continuous stochastic processes (i.e. the method of the conditional mathematical expectation rule) can not be applied directly.

Markov chain  $\xi(t)$  with continuous time and with the states  $1, 2, \dots, N$ , is completely described by the intensities  $\lambda_1, \dots, \lambda_N$  and by the matrix of the transfer probabilities  $P_{ij}(P_{ii} = 0)$  at the jumps moments. Time  $\eta_i$  of stay in state  $i$  has an exponential distribution with pdf  $p_{\eta_i} = \lambda_i \cdot \exp(-\lambda_i t)$ ,  $t > 0$ . Let us designate  $t_0, t_1, \dots, t_n, t_{n+1}$  as sampling moments. Let us  $\xi(t_n) = i$ . It is necessary to find the time interval  $T_i$  determining the next sampling moment  $t_{n+1} = t_n + T_i$  under the next conditions: 1) **condition of accuracy**: the variance  $V\hat{\tau}_{ij}$  of the estimation  $\hat{\tau}_{ij}$  of the jump moment  $\tau_{ij}$  from the state  $i$  into the state  $j$  ( $j \neq i$ ) is not more than a given value  $\sigma^2$  (the same for all  $i$  and  $j$ ); 2) **condition of miss**: probability of state miss on interval  $(t_n, t_n + T_i)$  is not more than a given value  $\gamma$ .

It is obtained conditional probability density for the jump moment under condition  $\{\xi(t_n) = i, \xi(t_n + T_i) = j\}$ :

$$p(t|i, j) = Ce^{-(\lambda_i - \lambda_j)t}, \quad 0 < t < T, \quad (1)$$

this is the cut exponential distribution (uniform distribution if  $\lambda_i = \lambda_j$ ),  $C$  is the normalizing constant.

Estimation  $\hat{\tau}_{ij}$  is the corresponding expectation

$$\hat{\tau}_{ij} = E\{\tau_{ij}|i, j\} = \begin{cases} [1 - \mu_{ij}T / (e^{\mu_{ij}T} - 1)] \mu_{ij}^{-1}, & \mu_{ij} \neq 0, \\ T/2, & \mu_{ij} = 0, \end{cases} \quad \mu_{ij} \equiv \lambda_i - \lambda_j \quad (2)$$

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Estimation variance is determined and interval  $T' \equiv T'_{i\sigma}$  is obtained from the accuracy condition

$$\max_{j, j \neq i, P_{ij} \neq 0} V \hat{\tau}_{ij} = \sigma^2,$$

it gives equation

$$\left[ 1 - (\mu_i^* T)^2 \left( e^{\mu_i^* T} + e^{-\mu_i^* T} - 2 \right)^{-1} \right] (\mu_i^*)^{-2} = \sigma^2 \quad (3)$$

where  $\mu_i^* = \min_{j, i \neq j, P_{ij} \neq 0} |\lambda_i - \lambda_j|$ . The graph way for determination of  $T \equiv T'_{i\sigma}$  is given. It is clear that the graph of transitions influences at the sampling procedure but not values  $P_{ij}$  of probability transitions.

Condition on probability of the state miss is reduced to the view

$$\max_{j, i \neq j, P_{ij} \neq 0} P\{\eta_i + \eta_j < T\} \equiv P\{\alpha(\eta_i + \eta_{j^*}) < \alpha T\} \equiv F_{\alpha\Sigma}(\alpha T) = \gamma, \quad (4)$$

where maximum is achieved on the state  $j^*$ . Distribution function of the sum  $\alpha\Sigma \equiv \alpha(\eta_i + \eta_{j^*})$  is depended from two parameters  $\lambda_i$  and  $\lambda_{j^*}$ ; it is reduced to the one-parameter family by substitution  $\alpha = \max(\lambda_i, \lambda_{j^*})$ ,  $\beta = \min(\lambda_i, \lambda_{j^*})$ ,  $k = \beta/\alpha$ ,  $0 < k \leq 1$ :

$$F_{\alpha\Sigma}(t) = \begin{cases} 1 - (e^{-kt}/k - e^{-t})k/(1-k), & k \neq 1, \\ 1 - (1+t)e^{-t}, & k = 1. \end{cases}$$

Approximate value is  $T_\gamma \approx \sqrt{2\gamma/(\lambda_i \lambda_{j^*})}$ ; Approximation accuracy is better than 10%. From conditions (3), (4) sampling interval in i state is  $T'_i = \min(T'_{i\sigma}, T'_{i\sigma})$ . Illustrate example is given.

## References

1. M. S Yarlykov and M. A. Mironov, *The Markov theory of Estimating Random processes*. New York Begell House, 1996.
2. V. A. Kazakov, One-dimensional kinetic equations for non-Markovian processes and statistical analysis problems of systems driven by the asymmetrical binary Markovian process, *Signal processing*, vol. 79, No. 1, pp. 87-95, 1999.